Week 11
Streaming Algorithms and the JL Lemma

Ryan Ziegler

## Outline

## Background <br> Probability <br> Streaming and Sketching Algorithms

Streaming $\ell_{2}$ Estimation

From Stream to Matrix

Conclusion

## Section 1

Background

Subsection 1

Probability

## A Probability Refresher

- (Discrete) probability distribution: given a set $S$ assign some probability $p_{i}$ to each element, so that $\sum p_{i}=1$

$$
\begin{aligned}
& S=\{R, G, B\} \\
& \mathbb{P}(R)=, 5 \\
& \mathbb{P}(G)=-25 \\
& \mathbb{P}(B)=\frac{.25}{1}
\end{aligned}
$$

## A Probability Refresher

- (Discrete) probability distribution: given a set $S$ assign some probability $p_{i}$ to each element, so that $\sum p_{i}=1$
- A random variable $X$ from a distribution $D$ is a variable whose value is randomly chosen according to some probability distribution $D$. Often denoted $X \sim D$.

$$
\begin{aligned}
& \mathbb{E}[X]=\Sigma_{i} p_{i} S_{i} \\
& \mathbb{E}[X+Y]=\mathbb{E}[X]+\mathbb{E}[Y] \\
& \operatorname{Var}(X)=\mathbb{E}\left((X-\mathbb{E} X)^{2}\right)=\mathbb{E}\left[X^{\partial}\right]-\mathbb{E}[X]^{2}
\end{aligned}
$$

## A Probability Refresher

- (Discrete) probability distribution: given a set $S$ assign some probability $p_{i}$ to each element, so that $\sum p_{i}=1$
- A random variable $X$ from a distribution $D$ is a variable whose value is randomly chosen according to some probability distribution $D$. Often denoted $X \sim D$.
- Expected value: suppose $S \subseteq \mathbb{R}$, then $\mathbb{E}[X]=\sum p_{i} S_{i}$. Intuitively, if we picked a bunch of $X$ following $D$, this is the average value we'd see.


## A Probability Refresher

- (Discrete) probability distribution: given a set $S$ assign some probability $p_{i}$ to each element, so that $\sum p_{i}=1$
- A random variable $X$ from a distribution $D$ is a variable whose value is randomly chosen according to some probability distribution $D$. Often denoted $X \sim D$.
- Expected value: suppose $S \subseteq \mathbb{R}$, then $\mathbb{E}[X]=\sum p_{i} S_{i}$. Intuitively, if we picked a bunch of $X$ following $D$, this is the average value we'd see.
- Expectation is a linear operator: $\mathbb{E}[X+Y]=\mathbb{E}[X]+\mathbb{E}[Y]$


## A Probability Refresher

- (Discrete) probability distribution: given a set $S$ assign some probability $p_{i}$ to each element, so that $\sum p_{i}=1$
- A random variable $X$ from a distribution $D$ is a variable whose value is randomly chosen according to some probability distribution $D$. Often denoted $X \sim D$.
- Expected value: suppose $S \subseteq \mathbb{R}$, then $\mathbb{E}[X]=\sum p_{i} S_{i}$. Intuitively, if we picked a bunch of $X$ following $D$, this is the average value we'd see.
- Expectation is a linear operator: $\mathbb{E}[X+Y]=\mathbb{E}[X]+\mathbb{E}[Y]$
- Variance: $\operatorname{Var}(X)=\mathbb{E}\left[X^{2}\right]-\mathbb{E}[X]^{2}$, a low variance indicates that most of the time, when we pick $X$ it will be close to $\mathbb{E}[X]$


## A Probability Refresher

- (Discrete) probability distribution: given a set $S$ assign some probability $p_{i}$ to each element, so that $\sum p_{i}=1$
- A random variable $X$ from a distribution $D$ is a variable whose value is randomly chosen according to some probability distribution $D$. Often denoted $X \sim D$.
- Expected value: suppose $S \subseteq \mathbb{R}$, then $\mathbb{E}[X]=\sum p_{i} S_{i}$. Intuitively, if we picked a bunch of $X$ following $D$, this is the average value we'd see.
- Expectation is a linear operator: $\mathbb{E}[X+Y]=\mathbb{E}[X]+\mathbb{E}[Y]$
- Variance: $\operatorname{Var}(X)=\mathbb{E}\left[X^{2}\right]-\mathbb{E}[X]^{2}$, a low variance indicates that most of the time, when we pick $X$ it will be close to $\mathbb{E}[X]$
- Note that for $c \in \mathbb{R}, \operatorname{Var}(c X)=c^{2} \operatorname{Var}(X)$

Even More Probability

$$
\mathcal{N}(0,1) \quad \mathbb{E}=0 \quad \operatorname{Var}=1
$$

- Normal distribution: $\mathcal{N}\left(\mu, \sigma^{2}\right)=\frac{1}{\sigma \sqrt{2 \pi}} \exp \left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^{2}\right)$

$X, Y \sim N(\mu, \sigma) X+Y^{\prime}$ is also Normal


## Even More Probability

- Normal distribution: $\mathcal{N}\left(\mu, \sigma^{2}\right)=\frac{1}{\sigma \sqrt{2 \pi}} \exp \left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^{2}\right)$
- Normal distribution is 2 -stable: for $X \sim \mathcal{N}\left(\mu_{1}, \sigma_{1}^{2}\right)$ and $Y \sim \mathcal{N}\left(\mu_{2}, \sigma_{2}^{2}\right), X+Y \sim \mathcal{N}\left(\mu_{1}+\mu_{2}, \sigma_{1}^{2}+\sigma_{2}^{2}\right)$


## Even More Probability

- Normal distribution: $\mathcal{N}\left(\mu, \sigma^{2}\right)=\frac{1}{\sigma \sqrt{2 \pi}} \exp \left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^{2}\right)$
- Normal distribution is 2-stable: for $X \sim \mathcal{N}\left(\mu_{1}, \sigma_{1}^{2}\right)$ and $Y \sim \mathcal{N}\left(\mu_{2}, \sigma_{2}^{2}\right), X+Y \sim \mathcal{N}\left(\mu_{1}+\mu_{2}, \sigma_{1}^{2}+\sigma_{2}^{2}\right)$
- $\chi^{2}(k)$ distribution: Sum of $k \mathcal{N}(0,1)$ random variables, has expected value $k$


## Even More Probability

- Normal distribution: $\mathcal{N}\left(\mu, \sigma^{2}\right)=\frac{1}{\sigma \sqrt{2 \pi}} \exp \left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^{2}\right)$
- Normal distribution is 2-stable: for $X \sim \mathcal{N}\left(\mu_{1}, \sigma_{1}^{2}\right)$ and $Y \sim \mathcal{N}\left(\mu_{2}, \sigma_{2}^{2}\right), X+Y \sim \mathcal{N}\left(\mu_{1}+\mu_{2}, \sigma_{1}^{2}+\sigma_{2}^{2}\right)$
- $\chi^{2}(k)$ distribution: Sum of $k \mathcal{N}(0,1)$ random variables, has expected value $k$
- Bernoulli distribution: If $X \sim \operatorname{Bernoulli}(p), X$ is 1 with probability $p$ and 0 with probability $(1-p)$


## Independence and Inequalities

$$
X_{1}, \ldots, X_{n}
$$

- A set of random variables is $k$-wise independent iff for any $k$ variables in the set, $f\left(x_{1}, \ldots, x_{k}\right)=f\left(x_{1}\right) \cdots f\left(x_{k}\right)$


## Independence and Inequalities

- A set of random variables is $k$-wise independent iff for any $k$ variables in the set, $f\left(x_{1}, \ldots, x_{k}\right)=f\left(x_{1}\right) \cdots f\left(x_{k}\right)$
- For $k$-wise independent random variables, $\mathbb{E}\left[\prod_{i=1}^{k} X_{i}\right]=\prod_{i=1}^{k} \mathbb{E}\left[x_{i}\right]$


## Independence and Inequalities

- A set of random variables is $k$-wise independent iff for any $k$ variables in the set, $f\left(x_{1}, \ldots, x_{k}\right)=f\left(x_{1}\right) \cdots f\left(x_{k}\right)$
- For $k$-wise independent random variables, $\mathbb{E}\left[\prod_{i=1}^{k} X_{i}\right]=\prod_{i=1}^{k} \mathbb{E}\left[x_{i}\right]$
- Important: $k$-wise independence implies $(k-1)$-wise independence


## Independence and Inequalities

- A set of random variables is $k$-wise independent iff for any $k$ variables in the set, $f\left(x_{1}, \ldots, x_{k}\right)=f\left(x_{1}\right) \cdots f\left(x_{k}\right)$
- For $k$-wise independent random variables, $\mathbb{E}\left[\prod_{i=1}^{k} X_{i}\right]=\prod_{i=1}^{k} \mathbb{E}\left[x_{i}\right]$
- Important: $k$-wise independence implies $(k-1)$-wise independence
- Chebyshev's inequality: $\left.\mathrm{P}(|X-\mathbb{E}[X]| \geq k \sigma) \leq \frac{1}{k^{2}}\right)$


## Independence and Inequalities

- A set of random variables is $k$-wise independent ff for any $k$ variables in the set, $f\left(x_{1}, \ldots, x_{k}\right)=f\left(x_{1}\right) \cdots f\left(x_{k}\right)$
- For $k$-wise independent random variables, $\mathbb{E}\left[\prod_{i=1}^{k} X_{i}\right]=\prod_{i=1}^{k} \mathbb{E}\left[x_{i}\right]$
- Important: $k$-wise independence implies $(k-1)$-wise independence
- Chebyshev's inequality: $\mathrm{P}(|X-\mathbb{E}[X]| \geq k \sigma) \leq \frac{1}{k^{2}}$
- Chernoff bound: Let $X$ be sum of $h$ fully independent Bernoulli RVs, and $\delta \geq 1$. $\mathrm{P}(X>(1+\delta) \mathbb{E}[X]) \leq e^{-\delta^{2} \mu / 3}$

$$
\begin{aligned}
Y_{1}, \ldots, Y_{n} \sim \operatorname{Bervolli}(p) \quad & X=\sum Y_{i} \\
& \mathbb{E}[X]=n p
\end{aligned}
$$

## Subsection 2

Streaming and Sketching Algorithms

Intro to Streaming Algorithms

- Streaming model: your algorithm receives inputs one-by-one, and you don't know how many inputs you'll receive. Too many inputs to store them all and compute later
too man elements to stere in memory!.
$\Rightarrow$ sublinear space
ex: you receive a stream of Youtube
Video Views, want the $k$ most watched videos today
keep a data structure with OIk) space, and update flat when sos see a woden


## Intro to Streaming Algorithms

- Streaming model: your algorithm receives inputs one-by-one, and you don't know how many inputs you'll receive. Too many inputs to store them all and compute later
- (*) Example: suppose you want to calculate the $k$ most watched YouTube videos today. It takes too much space to store all the YouTube videos and associated view counters, so you want an algorithm that does the following: upon recieving a YouTube video ID, update some data structure and continue without storing anything on disk. At the end of the day, this data structure should tell you the $k$ most viewed videos.


## Intro to Streaming Algorithms

- Streaming model: your algorithm receives inputs one-by-one, and you don't know how many inputs you'll receive. Too many inputs to store them all and compute later
- (*) Example: suppose you want to calculate the $k$ most watched YouTube videos today. It takes too much space to store all the YouTube videos and associated view counters, so you want an algorithm that does the following: upon recieving a YouTube video ID, update some data structure and continue without storing anything on disk. At the end of the day, this data structure should tell you the $k$ most viewed videos.
- (*) The above is possible to do exactly with only $O(k)$ space, but this is rare. Most streaming algorithms will only output approximates that are good with some probability
output a random variable $Z$

$$
\frac{\mathbb{E}[z]=g(\sigma)}{\operatorname{Lo}_{\sigma} \text { is our stream }}
$$

generally, $\operatorname{Var}(z) \leq g(\sigma)$
$Z^{*}=\frac{1}{\hbar} \sum Z_{l} \quad$ where $Z_{i}$ is an II copy of alg .

$$
\operatorname{Var}\left(z^{*}\right)=\frac{1}{n} \operatorname{Var}(z)
$$

$$
\mathbb{P}\left\{\mathbb{Z}^{*}-g(\sigma) \mid \geqslant \varepsilon g(\sigma)\right\} \leqslant \frac{1}{4} \Longrightarrow \text { (hebushev }
$$

$Q$ : how big shard $h$ be?

$$
h=O\left(\frac{1}{\varepsilon^{2}}\right) \text { if } \varepsilon \leqslant .001, h \leqslant 1000
$$

next goal: we want to output some $z^{\prime}$
so that

$$
\left.\mathbb{P} \xi\left|i^{-}-g(\sigma)\right| \geqslant 2 g(\sigma)\right\} \leqslant \delta
$$

$Z_{1}^{*}, \ldots, Z_{k}^{*}$ that fail $w /$ prob. $1 / 4$
$\downarrow$

$$
X_{1, \ldots}, X_{k} \quad \text { Bernolli(1/4) }
$$

$x_{i}=1$ if $Z_{i}$ is "bad", $O$ otherwise

median is bad of $n / 2$ bad values

$$
\begin{aligned}
& N=8 x_{i} \quad \mathbb{P}\left\{\mathbb{R}^{\prime}-g(\sigma)(\geqslant \varepsilon g(\sigma)\}\right.=\mathbb{P}\left\{N \geq \frac{n}{2}\right\} \\
&\left.O\left(\frac{\ln (1 / 8)}{a^{3}}\right) \text { pardlel upies }\right) \\
&=\mathbb{P}\left\{N \geqslant(1+1) \frac{n}{4}\right\} \\
& \leqslant \exp (-k / 12) \Rightarrow(6-O(61 / 8)
\end{aligned}
$$

## A Template for Sketching Algorithms

- First, output a random variable $Z$ such that $\mathbb{E}[Z]=g(\sigma)$ where $g(\sigma)$ is the function we're estimating for the stream $\sigma$
- Usually $Z$ will have high variance, typically $\operatorname{Var}(Z) \leq g(\sigma)$
- How to reduce variance? Run the streaming algorithm $h$ times in parallel, and let $Z^{*}=\frac{1}{h} \sum Z_{i}$

$$
\operatorname{Var}\left(Z^{*}\right)=\frac{1}{h} \operatorname{Var}\left(Z_{1}\right) \text { and } \mathbb{E}\left[Z^{*}\right]=\mathbb{E}\left[Z_{1}\right]
$$

- (*) By Chebyshev's inequality,

$$
\mathrm{P}\left(\left|Z^{*}-g(\sigma)\right|>\epsilon g(\sigma)\right) \leq \frac{\epsilon^{2}}{h}
$$

- (*) So, pick $h=\frac{4}{\epsilon^{2}}$ for constant failure probability of $\frac{1}{4}$


## The Median Trick

- Next goal: $\left|Z^{*}-g(\sigma)\right|>\epsilon g(\sigma)$ with some small probability $\delta$


## The Median Trick

- Next goal: $\left|Z^{*}-g(\sigma)\right|>\epsilon g(\sigma)$ with some small probability $\delta$
- Naive approach: do Chebyshev's again. Requires $O\left(\frac{1}{\delta \epsilon^{2}}\right)$ parallel copies. We want to do better


## The Median Trick

- Next goal: $\left|Z^{*}-g(\sigma)\right|>\epsilon g(\sigma)$ with some small probability $\delta$
- Naive approach: do Chebyshev's again. Requires $O\left(\frac{1}{\delta \epsilon^{2}}\right)$ parallel copies. We want to do better
- Consider parallel copies $Z_{1}^{*}, \ldots, Z_{k}^{*}$ that each fail with probability $1 / 4$


## The Median Trick

- Next goal: $\left|Z^{*}-g(\sigma)\right|>\epsilon g(\sigma)$ with some small probability $\delta$
- Naive approach: do Chebyshev's again. Requires $O\left(\frac{1}{\delta \epsilon^{2}}\right)$ parallel copies. We want to do better
- Consider parallel copies $Z_{1}^{*}, \ldots, Z_{k}^{*}$ that each fail with probability 1/4
- Our intuition tells us the median of these estimators should be "good" but how good?


## The Median Trick

- Next goal: $\left|Z^{*}-g(\sigma)\right|>\epsilon g(\sigma)$ with some small probability $\delta$
- Naive approach: do Chebyshev's again. Requires $O\left(\frac{1}{\delta \epsilon^{2}}\right)$ parallel copies. We want to do better
- Consider parallel copies $Z_{1}^{*}, \ldots, Z_{k}^{*}$ that each fail with probability 1/4
- Our intuition tells us the median of these estimators should be "good" but how good?
- (*) Let $X_{i}=1$ iff the $i$ th parallel copy fails, so then $X_{i} \sim \operatorname{Bernoulli}(1 / 4)$


## The Median Trick

- Next goal: $\left|Z^{*}-g(\sigma)\right|>\epsilon g(\sigma)$ with some small probability $\delta$
- Naive approach: do Chebyshev's again. Requires $O\left(\frac{1}{\delta \epsilon^{2}}\right)$ parallel copies. We want to do better
- Consider parallel copies $Z_{1}^{*}, \ldots, Z_{k}^{*}$ that each fail with probability 1/4
- Our intuition tells us the median of these estimators should be "good" but how good?
- (*) Let $X_{i}=1$ iff the $i$ th parallel copy fails, so then $X_{i} \sim \operatorname{Bernoulli}(1 / 4)$
- (*) Define $X=\sum X_{i}$, so then $\mathbb{E}[X]=\frac{k}{4}$


## The Median Trick

- Next goal: $\left|Z^{*}-g(\sigma)\right|>\epsilon g(\sigma)$ with some small probability $\delta$
- Naive approach: do Chebyshev's again. Requires $O\left(\frac{1}{\delta \epsilon^{2}}\right)$ parallel copies. We want to do better
- Consider parallel copies $Z_{1}^{*}, \ldots, Z_{k}^{*}$ that each fail with probability 1/4
- Our intuition tells us the median of these estimators should be "good" but how good?
- (*) Let $X_{i}=1$ iff the $i$ th parallel copy fails, so then $X_{i} \sim \operatorname{Bernoulli}(1 / 4)$
- (*) Define $X=\sum X_{i}$, so then $\mathbb{E}[X]=\frac{k}{4}$
- (*) By Chernoff bound,

$$
\mathrm{P}\left(X \geq(1+1) \frac{k}{4}\right) \leq e^{-k / 12}
$$

## The Median Trick

- Next goal: $\left|Z^{*}-g(\sigma)\right|>\epsilon g(\sigma)$ with some small probability $\delta$
- Naive approach: do Chebyshev's again. Requires $O\left(\frac{1}{\delta \epsilon^{2}}\right)$ parallel copies. We want to do better
- Consider parallel copies $Z_{1}^{*}, \ldots, Z_{k}^{*}$ that each fail with probability 1/4
- Our intuition tells us the median of these estimators should be "good" but how good?
- (*) Let $X_{i}=1$ iff the $i$ th parallel copy fails, so then $X_{i} \sim \operatorname{Bernoulli}(1 / 4)$
- (*) Define $X=\sum X_{i}$, so then $\mathbb{E}[X]=\frac{k}{4}$
- (*) By Chernoff bound,

$$
\mathrm{P}\left(X \geq(1+1) \frac{k}{4}\right) \leq e^{-k / 12}
$$

- $(*)$ So, pick $k=O(\log (1 / \delta))$. Only running $O\left(\frac{\log \left(\frac{1}{\delta}\right)}{\epsilon^{2}}\right)$

Section 2

Streaming $\ell_{2}$ Estimation

## Frequency Moment Estimation

- Problem: we receive a stream $\sigma$ of values $e_{1}, \cdots \in \mathbb{Z}$ where $1 \leq e_{i} \leq n$ for some $n$ we know apriori


## Frequency Moment Estimation

- Problem: we receive a stream $\sigma$ of values $e_{1}, \cdots \in \mathbb{Z}$ where $1 \leq e_{i} \leq n$ for some $n$ we know apriori
- Define the frequency vector to be $f(\sigma)=\left(f_{1}, \ldots, f_{n}\right)$ where $f_{i}$ is the number of times we've seen $i$

$$
\begin{aligned}
& \sigma=\{1,1,5,7\} \\
& f=(2,0,0,0,1,0,1)
\end{aligned}
$$

## Frequency Moment Estimation

- Problem: we receive a stream $\sigma$ of values $e_{1}, \cdots \in \mathbb{Z}$ where $1 \leq e_{i} \leq n$ for some $n$ we know apriori
- Define the frequency vector to be $f(\sigma)=\left(f_{1}, \ldots, f_{n}\right)$ where $f_{i}$ is the number of times we've seen $i$
- Goal: estimate $\|f(\sigma)\|_{2}^{2}$ with only $O(\operatorname{poly} \log (n))$ space


## Frequency Moment Estimation

- Problem: we receive a stream $\sigma$ of values $e_{1}, \cdots \in \mathbb{Z}$ where $1 \leq e_{i} \leq n$ for some $n$ we know apriori
- Define the frequency vector to be $f(\sigma)=\left(f_{1}, \ldots, f_{n}\right)$ where $f_{i}$ is the number of times we've seen $i$
- Goal: estimate $\|f(\sigma)\|_{2}^{2}$ with only $O(\operatorname{poly} \log (n))$ space
- Recall the definition of $L_{2}$ norm:

$$
\|f(\sigma)\|_{2}^{2}=\sum_{i=1}^{n} f_{i}^{2}
$$

AMS F2 Estimation

- Intuition: keep a single variable $Z$ so that we can output $Z^{2}$ as our estimate of $\|f(\sigma)\|_{2}^{2}$

$$
\mathbb{E}\left[2^{4}\right] \leqslant\|f(\sigma)\|_{\partial}^{\partial}
$$

$$
\begin{aligned}
Z= & \sum f_{i} Y_{i} \\
& \mathbb{E}[z]=\sum f_{i} \mathbb{E}\left[Y_{i}\right] \\
& \mathbb{E}\left[z^{2}\right]=\underbrace{\sum f_{i}^{2} \mathbb{E}\left[Y_{i}\right]^{2}}_{\mathbb{E}}+2 Y_{i \neq j} \mathbb{E}\left[Y_{i} Y_{j}\right] f_{i} f_{j} \\
& \mathbb{E}\left[Y_{i}^{2}\right]=\mathbb{I} \quad \mathbb{E}\left[Z^{2}\right]=\|f(\sigma)\|_{2}^{2}
\end{aligned}
$$

let $Y_{i}$ be pa ind $Y_{i}=1$ awl prob $\cdot 5$ $Y_{i}=-1 \quad \omega 1$ prob -5

## AMS F2 Estimation

- Intuition: keep a single variable $Z$ so that we can output $Z^{2}$ as our estimate of $\|f(\sigma)\|_{2}^{2}$
- (*) Idea: create some random variable $Y_{i}$ for each index so that $\mathbb{E}\left[Z^{2}\right]=\|f(\sigma)\|_{2}^{2}$. In particular, $Z=\sum Y_{i} f_{i}$

$$
\mathbb{E}\left[Z^{2}\right]=\sum f_{i}^{2} Y_{i}^{2}+2 \sum_{i \neq j} f_{i} f_{j} Y_{i} Y_{j}
$$

- (*) We need $Y_{i}$ to be pairwise independent and satisfy $\mathbb{E}\left[Y_{i} Y_{j}\right]=0$ and $\mathbb{E}\left[Y_{i}^{2}\right]=1$


## AMS F2 Estimation

- Intuition: keep a single variable $Z$ so that we can output $Z^{2}$ as our estimate of $\|f(\sigma)\|_{2}^{2}$
- $\left.{ }^{*}\right)$ Idea: create some random variable $Y_{i}$ for each index so that $\mathbb{E}\left[Z^{2}\right]=\|f(\sigma)\|_{2}^{2}$. In particular, $Z=\sum Y_{i} f_{i}$

$$
\mathbb{E}\left[Z^{2}\right]=\sum f_{i}^{2} Y_{i}^{2}+2 \sum_{i \neq j} f_{i} f_{j} Y_{i} Y_{j}
$$

- (*) We need $Y_{i}$ to be pairwise independent and satisfy $\mathbb{E}\left[Y_{i} Y_{j}\right]=0$ and $\mathbb{E}\left[Y_{i}^{2}\right]=1$
- (*) Solution: $Y_{i}=1$ with probability $\frac{1}{2}$ and $Y_{i}=-1$ with probability $\frac{1}{2}$

AMS F2 Estimation Continued

- Creating $O(n)$ random variables takes up too much space!
- Solution: $O(1)$-wise independent hash family of functions $[n] \rightarrow\{-1,1\}$ can be stored in $O(\operatorname{polylog}(n))$ space

$$
z=\Sigma f_{c} h(i)
$$

when we see $e \in(n)$

$$
z^{x}=h(e)
$$

## AMS F2 Estimation Continued

- Creating $O(n)$ random variables takes up too much space!
- Solution: $O(1)$-wise independent hash family of functions $[n] \rightarrow\{-1,1\}$ can be stored in $O(\operatorname{polylog}(n))$ space
- (*) Replace each $Y_{i}$ with $h(i)$, and the analysis is the exact same


## AMS F2 Estimation Continued

- Creating $O(n)$ random variables takes up too much space!
- Solution: $O(1)$-wise independent hash family of functions $[n] \rightarrow\{-1,1\}$ can be stored in $O(\operatorname{polylog}(n))$ space
- $\left.{ }^{*}\right)$ Replace each $Y_{i}$ with $h(i)$, and the analysis is the exact same
- (*) Similar analysis shows $\mathbb{E}\left[Z^{4}\right] \leq 2\|f(\sigma)\|_{2}^{2}$, so we can apply average and median idea from before


## AMS F2 Estimation Continued

- Creating $O(n)$ random variables takes up too much space!
- Solution: $O(1)$-wise independent hash family of functions $[n] \rightarrow\{-1,1\}$ can be stored in $O(\operatorname{polylog}(n))$ space
- (*) Replace each $Y_{i}$ with $h(i)$, and the analysis is the exact same
- (*) Similar analysis shows $\mathbb{E}\left[Z^{4}\right] \leq 2\|f(\sigma)\|_{2}^{2}$, so we can apply average and median idea from before

```
def ams_f2:
    let h be a hash function from hash family H
    let z = 0
    while i is an item from stream
        z = z + h(i)
    output z
```


## Extending F2 Estimation

- Note that we never used the fact that $f_{i}$ was positive or integral
- Richer model: receive a stream of updates of the form $\left(i, \Delta_{i}\right)$ representing a change to the $i$ th coordinate of our vector


## Extending F2 Estimation

- Note that we never used the fact that $f_{i}$ was positive or integral
- Richer model: receive a stream of updates of the form $\left(i, \Delta_{i}\right)$ representing a change to the $i$ th coordinate of our vector


## def l2_estimate:

let $h$ be a hash function from hash family $H$
let $\mathrm{z}=0$
while (id) is an item from stream
$z=z+h(i) d$
output $z^{2}$
alg is a fond $C$ let $\sigma_{1}$ and $\sigma_{\alpha}$ be streamer

$$
C\left(\sigma_{1}+\sigma_{2}\right)=C\left(\sigma_{1}\right)+C\left(\sigma_{2}\right)
$$

$$
O\left(\frac { ( n / 6 ) } { n ^ { 2 } } \left\{\left[\begin{array}{cccc}
\frac{n}{n_{2}^{(1)}} n_{2}(2) & \cdots & \cdots & n_{f}(n) \\
& \vdots & & \\
n_{v}(1) & \cdots & & \\
n_{6}(n)
\end{array}\right]\left[\begin{array}{l}
f
\end{array}\right]=\left[0\left(\frac{1 n^{\prime} / k_{0}}{a^{2}}\right)\right.\right.\right.
$$

Section 3
From Stream to Matrix

## Linear Sketching

- What we just created is a linear sketch: call our algorithm $C$. We can show that $C\left(\sigma_{1}+\sigma_{2}\right)=C\left(\sigma_{1}\right)+C\left(\sigma_{2}\right)$, since each iteration we add to $Z$
- $\left({ }^{*}\right)$ Geometric interpretation: our algorithm is an $\frac{\log (1 / \delta) \log n}{\epsilon^{2}} \times n$ matrix $M$ of $\{-1,1\}$ values, each row is a parallel copy of the streaming algorithm


## Linear Sketching

- What we just created is a linear sketch: call our algorithm $C$. We can show that $C\left(\sigma_{1}+\sigma_{2}\right)=C\left(\sigma_{1}\right)+C\left(\sigma_{2}\right)$, since each iteration we add to $Z$
- $(*)$ Geometric interpretation: our algorithm is an $\frac{\log (1 / \delta) \log n}{\epsilon^{2}} \times n$ matrix $M$ of $\{-1,1\}$ values, each row is a parallel copy of the streaming algorithm
- (*) Now we have $M x=y$ where $y$ is a vector whose length is similar to that of $x$ but is in lower dimension


## Linear Sketching

- What we just created is a linear sketch: call our algorithm $C$. We can show that $C\left(\sigma_{1}+\sigma_{2}\right)=C\left(\sigma_{1}\right)+C\left(\sigma_{2}\right)$, since each iteration we add to $Z$
- $(*)$ Geometric interpretation: our algorithm is an $\frac{\log (1 / \delta) \log n}{\epsilon^{2}} \times n$ matrix $M$ of $\{-1,1\}$ values, each row is a parallel copy of the streaming algorithm
- (*) Now we have $M x=y$ where $y$ is a vector whose length is similar to that of $x$ but is in lower dimension
- $\left(^{*}\right)$ Next goal: generalize this idea so that we can reduce the dimension of a set of vectors while preserving pairwise distances


## Linear Sketching

- What we just created is a linear sketch: call our algorithm $C$. We can show that $C\left(\sigma_{1}+\sigma_{2}\right)=C\left(\sigma_{1}\right)+C\left(\sigma_{2}\right)$, since each iteration we add to $Z$
- $(*)$ Geometric interpretation: our algorithm is an $\frac{\log (1 / \delta) \log n}{\epsilon^{2}} \times n$ matrix $M$ of $\{-1,1\}$ values, each row is a parallel copy of the streaming algorithm
- (*) Now we have $M x=y$ where $y$ is a vector whose length is similar to that of $x$ but is in lower dimension
- $\left(^{*}\right)$ Next goal: generalize this idea so that we can reduce the dimension of a set of vectors while preserving pairwise distances
- $\left(^{*}\right)$ Useful in real-world applications such as nearest neighbors, ML, etc

The JL Lemma

$$
\mathbb{E}[X]=0 \quad \mathbb{E}\left[x^{2}\right]=1
$$

- Let $M$ be an $k \times n$ matrix where each entry is chosen independently from $\mathcal{N}(0,1)$
- Claim: for $k=\Omega\left(\frac{\log (1 / \delta)}{\epsilon^{2}}\right)$, we have that with probability $1-\delta$, $\left\|\frac{1}{\sqrt{k}} M x\right\|_{2}=(1 \pm \epsilon)\|x\|_{2}$ for fixed $x \in \mathbb{R}^{n}$
Let $S$ be a set of $n$ vectors,
If $\delta=1 / n$ el $J L$ matrix will preserve pairwise distances


## The JL Lemma

- Let $M$ be an $k \times n$ matrix where each entry is chosen independently from $\mathcal{N}(0,1)$
- Claim: for $k=\Omega\left(\frac{\log (1 / \delta)}{\epsilon^{2}}\right)$, we have that with probability $1-\delta$, $\left\|\frac{1}{\sqrt{k}} M x\right\|_{2}=(1 \pm \epsilon)\|x\|_{2}$ for fixed $x \in \mathbb{R}^{n}$
- Immediate corollary: Let $S$ be a set of $k$ vectors in $\mathbb{R}^{n}$, we can preserve pairwise distances with high probability by picking $k=\Omega\left(\frac{\log n}{\epsilon^{2}}\right)$


## The JL Lemma

- Let $M$ be an $k \times n$ matrix where each entry is chosen independently from $\mathcal{N}(0,1)$
- Claim: for $k=\Omega\left(\frac{\log (1 / \delta)}{\epsilon^{2}}\right)$, we have that with probability $1-\delta$, $\left\|\frac{1}{\sqrt{k}} M x\right\|_{2}=(1 \pm \epsilon)\|x\|_{2}$ for fixed $x \in \mathbb{R}^{n}$
- Immediate corollary: Let $S$ be a set of $k$ vectors in $\mathbb{R}^{n}$, we can preserve pairwise distances with high probability by picking $k=\Omega\left(\frac{\log n}{\epsilon^{2}}\right)$


## JL Lemma: Idea of Proof

- Fix some vector $x$ (wlog, let $\|x\|=1$ ) and use 2-stability of Normal distribution


## JL Lemma: Idea of Proof

- Fix some vector $x$ (wlog, let $\|x\|=1$ ) and use 2-stability of Normal distribution
- $\left({ }^{*}\right)$ Let $y=M x$, so then $y_{i}=\sum_{j=1}^{k} M_{i j} x_{i}$


## JL Lemma: Idea of Proof

- Fix some vector $x$ (wlog, let $\|x\|=1$ ) and use 2-stability of Normal distribution
- (*) Let $y=M x$, so then $y_{i}=\sum_{j=1}^{k} M_{i j} x_{i}$
- (*) $y$ is a Normal vector in $\mathbb{R}^{k}$, and each $y_{i}$ is $\mathcal{N}(0,1)$ (variance because $\sum x_{i}^{2}=1$ )


## JL Lemma: Idea of Proof

- Fix some vector $x$ (wlog, let $\|x\|=1$ ) and use 2-stability of Normal distribution
- (*) Let $y=M x$, so then $y_{i}=\sum_{j=1}^{k} M_{i j} x_{i}$
- (*) $y$ is a Normal vector in $\mathbb{R}^{k}$, and each $y_{i}$ is $\mathcal{N}(0,1)$ (variance because $\sum x_{i}^{2}=1$ )
- (*) Let $\alpha=\sum y_{i}^{2}$, so then $\alpha \sim \chi^{2}(k)$


## JL Lemma: Idea of Proof

- Fix some vector $x$ (wlog, let $\|x\|=1$ ) and use 2-stability of Normal distribution
- (*) Let $y=M x$, so then $y_{i}=\sum_{j=1}^{k} M_{i j} x_{i}$
- (*) $y$ is a Normal vector in $\mathbb{R}^{k}$, and each $y_{i}$ is $\mathcal{N}(0,1)$ (variance because $\sum x_{i}^{2}=1$ )
- (*) Let $\alpha=\sum y_{i}^{2}$, so then $\alpha \sim \chi^{2}(k)$
- (*) Thus $\mathrm{P}\left((1-\epsilon)^{2} k \leq \alpha \leq(1+\epsilon)^{2} k\right) \geq 1-2 e^{O(1) \epsilon^{2} k}$


## JL Lemma: Idea of Proof

- Fix some vector $x$ (wlog, let $\|x\|=1$ ) and use 2-stability of Normal distribution
- (*) Let $y=M x$, so then $y_{i}=\sum_{j=1}^{k} M_{i j} x_{i}$
- (*) $y$ is a Normal vector in $\mathbb{R}^{k}$, and each $y_{i}$ is $\mathcal{N}(0,1)$ (variance because $\sum x_{i}^{2}=1$ )
- (*) Let $\alpha=\sum y_{i}^{2}$, so then $\alpha \sim \chi^{2}(k)$
- $\left.{ }^{*}\right)$ Thus $\mathrm{P}\left((1-\epsilon)^{2} k \leq \alpha \leq(1+\epsilon)^{2} k\right) \geq 1-2 e^{O(1) \epsilon^{2} k}$
- (*) Picking $k=\Omega\left(\frac{\log (1 / \delta)}{\epsilon^{2}}\right)$ gets us the probability we want

Section 4
Conclusion

## JL Lemma: Intuition and Application

- Why does projecting to a random subspace work? A large enough random subspace means errors induced by "bad vectors" (i.e. those orthogonal to many rows in the matrix) have extremely low probability of ocurring


## JL Lemma: Intuition and Application

- Why does projecting to a random subspace work? A large enough random subspace means errors induced by "bad vectors" (i.e. those orthogonal to many rows in the matrix) have extremely low probability of ocurring
- Useful for tasks such as clustering/ML: things closer together/more similar in low dimension will be close in high dimension, so can reduce dimension and speed up clustering


## JL Lemma: Intuition and Application

- Why does projecting to a random subspace work? A large enough random subspace means errors induced by "bad vectors" (i.e. those orthogonal to many rows in the matrix) have extremely low probability of ocurring
- Useful for tasks such as clustering/ML: things closer together/more similar in low dimension will be close in high dimension, so can reduce dimension and speed up clustering
- Coreset generation: Many hard geometric problems have fast approximate solutions via coreset technique, which generates a set $S^{\prime}$ from input $S$ so that running an exact algorithm on $S^{\prime}$ generates a high accuracy approximation for that algorithm on $S$. JL technique can be used in generating coresets


## JL Lemma: Intuition and Application

- Why does projecting to a random subspace work? A large enough random subspace means errors induced by "bad vectors" (i.e. those orthogonal to many rows in the matrix) have extremely low probability of ocurring
- Useful for tasks such as clustering/ML: things closer together/more similar in low dimension will be close in high dimension, so can reduce dimension and speed up clustering
- Coreset generation: Many hard geometric problems have fast approximate solutions via coreset technique, which generates a set $S^{\prime}$ from input $S$ so that running an exact algorithm on $S^{\prime}$ generates a high accuracy approximation for that algorithm on $S$. JL technique can be used in generating coresets
- Key advantage of JL is that it is oblivious to data


## One more thing. . .

- JL Lemma extends to preserving vector distances in entire subspaces of $\mathbb{R}^{n}$ !


## One more thing. . .

- JL Lemma extends to preserving vector distances in entire subspaces of $\mathbb{R}^{n}$ !
- Let $E$ be a linear subspace of dimension $d$


## One more thing. . .

- JL Lemma extends to preserving vector distances in entire subspaces of $\mathbb{R}^{n}$ !
- Let $E$ be a linear subspace of dimension $d$
- Can preserve distances between vectors in $E$ with $k=\Omega\left(\frac{d \log (1 / \delta)}{\epsilon^{2}}\right)$


## One more thing. . .

- JL Lemma extends to preserving vector distances in entire subspaces of $\mathbb{R}^{n}$ !
- Let $E$ be a linear subspace of dimension $d$
- Can preserve distances between vectors in $E$ with $k=\Omega\left(\frac{d \log (1 / \delta)}{\epsilon^{2}}\right)$
- Works for all vectors in $E$, even though there are infinitely many!


## One more thing. . .

- JL Lemma extends to preserving vector distances in entire subspaces of $\mathbb{R}^{n}$ !
- Let $E$ be a linear subspace of dimension $d$
- Can preserve distances between vectors in $E$ with $k=\Omega\left(\frac{d \log (1 / \delta)}{\epsilon^{2}}\right)$
- Works for all vectors in $E$, even though there are infinitely many!
- Poof: consider partitioning the $d$ dimensional unit ball into small hypercubes with small side length. Show that preserving lengths of vectors to these hypercubes is sufficient to preserve lengths of all vectors.

$$
\hat{\xi} g_{i}\left(f_{i}\right)
$$



$$
0<k \leqslant 2
$$

